

Prediction

Table of contents

Prediction	1
Bivariate Normal Distribution	1
Prediction via Minimum MSE	3
Best Linear Predictor	4
Specialization to the Bivariate Normal	4
R Example	5

Prediction

Bivariate Normal Distribution

Let (X, Y) be bivariate normal.

$$f_{X,Y}(x, y) = \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)}\left[\left(\frac{x-\mu_X}{\sigma_X}\right)^2 + \left(\frac{y-\mu_Y}{\sigma_Y}\right)^2 - 2\rho\left(\frac{x-\mu_X}{\sigma_X}\right)\left(\frac{y-\mu_Y}{\sigma_Y}\right)\right]\right\}.$$

$-\infty < \mu_X < \infty$ and $-\infty < \mu_Y < \infty$.

$0 < \sigma_X, \sigma_Y$ and $-1 < \rho < 1$.

Marginals

$$X \sim \mathcal{N}(\mu_X, \sigma_X^2)$$

$$Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2).$$

Exercise: Show

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma_X} \exp \left\{ -\frac{1}{2} \left(\frac{x - \mu_X}{\sigma_X} \right)^2 \right\}$$

Exercise: Show

$$f_Y(y) = \frac{1}{\sqrt{2\pi}\sigma_Y} \exp \left\{ -\frac{1}{2} \left(\frac{y - \mu_Y}{\sigma_Y} \right)^2 \right\}.$$

Conditionals

$$Y \mid X = x \sim \mathcal{N} \left(\mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x - \mu_X), \sigma_Y^2 (1 - \rho^2) \right).$$

$$f_{Y|X}(y \mid x) = \frac{1}{\sigma_Y \sqrt{1 - \rho^2}} \exp \left\{ -\frac{1}{2} \frac{(y - \mu_Y - \rho \frac{\sigma_Y}{\sigma_X} (x - \mu_X))^2}{\sigma_Y^2 (1 - \rho^2)} \right\}.$$

$$\mathbb{E}[Y \mid X = x] = \mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x - \mu_X).$$

$$\text{Var}(Y \mid X = x) = \sigma_Y^2 (1 - \rho^2).$$

Covariance and Correlation

$$\text{Cov}(X, Y) = \rho \sigma_X \sigma_Y.$$

$$\text{Corr}(X, Y) = \rho.$$

Note: For a bivariate normal, X and Y are independent if and only if $\rho = 0$. \$\$

Prediction via Minimum MSE

A random variable is observed and then based on the observed value, we predict a second random variable.

Let $g(X)$ denote the predictor function.

Choose a function $g(X)$ so it tends to be close to Y

Criteria:

Goal: minimizing MSE, i.e., minimize

$$E[(Y - g(X))^2].$$

Claim: The best predictor is

$$g(x) = E[Y \mid X = x].$$

Proof:

Condition on X and then take $E[\]$.

$$E[(Y - g(X))^2 \mid X] = E[(Y - E[Y \mid X] + E[Y \mid X] - g(X))^2 \mid X]$$

$$= E[(Y - E[Y \mid X])^2 \mid X] + E[(E[Y \mid X] - g(X))^2 \mid X] + 2E[(Y - E[Y \mid X])(E[Y \mid X] - g(X)) \mid X].$$

Note: $E[Y \mid X] - g(X)$ given X can be treated as a constant.

$$E[Y - E[Y \mid X] \mid X] = 0 \Rightarrow \text{cross term} = 0.$$

$$\Rightarrow E[(Y - g(X))^2 \mid X] \geq E[(Y - E[Y \mid X])^2 \mid X].$$

$$\Rightarrow E[(Y - g(X))^2] \geq E[(Y - E[Y \mid X])^2].$$

Exercise: Show

$$E[(X - a)^2] \text{ is minimized at } a = E[X].$$

Best Linear Predictor

It sometimes happens that the joint density of X and Y is not completely known, or $E[Y|X]$ is difficult to calculate. However, $\mu_X, \mu_Y, \sigma_X^2, \sigma_Y^2, \sigma_{XY}$ are known, then we can at least determine the **Best Linear Predictor** of Y on X .

Assume

$$g(X) = a + bX \text{ and minimize } E[(Y - (a + bX))^2].$$

Answer:

$$\frac{\partial}{\partial a} E[(Y - (a + bX))^2] = -2 E[Y] + 2a + 2b E[X] = 0.$$

$$\frac{\partial}{\partial b} E[(Y - (a + bX))^2] = -2 E[XY] + 2a E[X] + 2b E[X^2] = 0.$$

$$b = \frac{E[XY] - E[X]E[Y]}{E[X^2] - (E[X])^2} = \frac{\text{Cov}(X, Y)}{\text{Var}(X)} = \rho \frac{\sigma_Y}{\sigma_X}.$$

$$a = E[Y] - b E[X] = \mu_Y - \rho \frac{\sigma_Y}{\sigma_X} \mu_X.$$

$$g(x) = a + bx = \mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x - \mu_X).$$

$$(\text{intercept} = \mu_Y - \rho \frac{\sigma_Y}{\sigma_X} \mu_X, \text{ slope} = \rho \frac{\sigma_Y}{\sigma_X}).$$

Specialization to the Bivariate Normal

$$g(x) = E[Y | X = x] = \mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x - \mu_X).$$

$$\text{MSE}(g) = E[(Y - \mu_Y - \rho \frac{\sigma_Y}{\sigma_X} (X - \mu_X))^2] = \sigma_Y^2 (1 - \rho^2).$$

Note: For (X, Y) Bivariate Normal, the Best Predictor is the Best Linear Predictor $E[Y|X]$.

Note: For (X, Y) Bivariate Normal, the Best Linear Predictor has MSE $V(Y|X)$.

Note: Thus, higher ρ , then lower prediction error.

R Example

```
set.seed(1)
n <- 2000
rho <- 0.6
muX <- 0; muY <- 1
sigX <- 2; sigY <- 3
Sigma <- matrix(c(sigX^2, rho*sigX*sigY,
                  rho*sigX*sigY, sigY^2), 2, 2)
Z <- MASS::mvrnorm(n, mu = c(muX, muY), Sigma = Sigma)
X <- Z[,1]; Y <- Z[,2]

b <- rho * sigY / sigX
a <- muY - b * muX
gX <- a + b * X

c(mean_mse = mean((Y - gX)^2),
  theory = sigY^2 * (1 - rho^2))
```

```
mean_mse  theory
6.372515 5.760000
```