

# Homework 3

1/15

## Example 8.1.1 election polling.

prior  $\pi \sim \text{Beta}(a, b)$   $p(\pi) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \pi^{a-1} (1-\pi)^{b-1}$

likelihood  $x_1, \dots, x_n \stackrel{\text{iid}}{\sim} \text{Ber}(\pi)$

p. 93  $a = 330$   
 $b = 270$

$X = \sum_{i=1}^n X_i \sim \text{Ber}(n, \pi)$

$p(x|\pi) = \binom{n}{x} \pi^x (1-\pi)^{n-x}$

p. 96 data

$x = 620$

$n = 1000$

posterior  $p(\pi|x) = \frac{p(\pi) p(x|\pi)}{\int p(\pi) p(x|\pi) d\pi}$

$p(\pi|x) \propto p(\pi) p(x|\pi)$

$\propto \pi^{a-1} (1-\pi)^{b-1} \pi^x (1-\pi)^{n-x}$

$\propto \pi^{a+x-1} (1-\pi)^{b+n-x-1}$

$\therefore \pi|x \sim \text{Beta}(a+x, b+n-x)$

posterior mean  $E[\pi|x] = \int_0^1 \pi p(\pi|x) d\pi$

$= \frac{a+x}{a+b+n}$

$= \left(\frac{a}{a+b}\right) \left[\frac{a+b}{a+b+n}\right] + \left(\frac{x}{n}\right) \left[\frac{n}{a+b+n}\right]$

p. 96  $E[\pi|x] = \frac{330+620}{330+270+1000} = \frac{950}{1600} = 0.59375$

Example 8.1.1 Beta-Binomial Election polling.

data

list(X = 620, n = 1000) # frequentist estimate 0.620

inits

list(p = 0.25) # starting values for p

model

model;

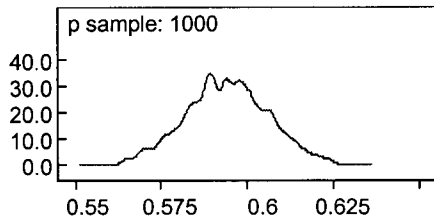
```

{
  X ~ dbin(p,n)
  p ~ dbeta(330,270) # prior mode = 0.55, 95% prior interval (0.51, 0.59)
}

```

Output:

*posterior density of p*



node	mean	sd	MC error	2.5%	median	97.5%	start	sample
p	0.5941	0.0123	4.757E-4	0.5699	0.5939	0.6182	1001	1000

*1*  
*posterior*  
*estimate*  
*of p*

*2*  
*posterior 95%*  
*credible*  
*interval for p*

Example B.1.2 Weighing an object.

prior  $\mu \sim N(\mu_0, \sigma_0^2)$

p. 93  $\mu_0 = 200$   
 $\sigma_0 = 10$

$$p(\mu) \propto \exp\left[-\frac{1}{2\sigma_0^2}(\mu - \mu_0)^2\right]$$

likelihood  $x_1, \dots, x_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$   $\sigma^2 = 1$  known.

$$p(x_1, \dots, x_n | \mu) \propto \exp\left[-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2\right]$$

p. 97  
data  
 $n=5$   
 $\bar{x} = 197.79$

posterior  $p(\mu | x_1, \dots, x_n) \propto p(\mu) p(x_1, \dots, x_n | \mu)$

$$\propto \exp\left[-\frac{1}{2\sigma_0^2}(\mu - \mu_0)^2\right] \cdot \exp\left[-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2\right]$$

$$\propto \exp\left[-\frac{1}{2} \left\{ \frac{(\mu - \mu_0)^2}{\sigma_0^2} + \sum_{i=1}^n \frac{(x_i - \mu)^2}{\sigma^2} \right\}\right]$$

complete the square

$$\propto \exp \left[ \frac{1}{2} \left( \frac{\mu^2}{\sigma_0^2} - \frac{2\mu\mu_0}{\sigma_0^2} + \frac{\mu_0^2}{\sigma_0^2} + \frac{\sum x_i^2}{\sigma^2} - \frac{2\mu \sum x_i}{\sigma^2} + \frac{n\mu^2}{\sigma^2} \right) \right]$$

$$\propto \exp \left[ \frac{1}{2} \left( \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right) \mu^2 - 2 \left( \frac{\mu_0}{\sigma_0^2} + \frac{\sum x_i}{\sigma^2} \right) \mu + \left( \frac{\mu_0^2}{\sigma_0^2} + \frac{\sum x_i^2}{\sigma^2} \right) \right) \right]$$

constant

$$\propto \exp \left[ -\frac{1}{2} \left( \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right) \mu^2 - 2 \left( \frac{\mu_0}{\sigma_0^2} + \frac{\sum x_i}{\sigma^2} \right) \mu \right) \right]$$

$$\propto \exp \left[ -\frac{1}{2} \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right) \left( \mu^2 - 2 \frac{\left( \frac{\mu_0}{\sigma_0^2} + \frac{\sum x_i}{\sigma^2} \right)}{\left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right)} \mu + k - k \right) \right]$$

let  $k = \frac{\left( \frac{\mu_0}{\sigma_0^2} + \frac{\sum x_i}{\sigma^2} \right)}{\left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right)}$

to complete the square add a constant, the  $-k$  goes away with  $\propto$

$$\propto \exp \left[ -\frac{1}{2} \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right) \left( \mu - \frac{\left( \frac{\mu_0}{\sigma_0^2} + \frac{\sum x_i}{\sigma^2} \right)}{\left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right)} \right)^2 \right]$$

$$\therefore \mu | x_1, \dots, x_n \sim N(\mu_n, \sigma_n^2)$$

posterior mean.

$$\mu_n = \frac{\left( \frac{\mu_0}{\sigma_0^2} + \frac{n\bar{x}}{\sigma^2} \right)}{\left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right)} = \mu_0 \left( \frac{\sigma_n^2}{\sigma_0^2} \right) + \bar{x} \left( \frac{\sigma_n^2}{\sigma^2} \right)$$

posterior variance.

$$\sigma_n^2 = \frac{1}{\left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right)}$$

$$\hat{\mu}_n = \frac{\left( \frac{200}{10^2} + \frac{(5)(197.79)}{1} \right)}{\frac{1}{10^2} + \frac{5}{1^2}} = \frac{990.95}{5.01} = 197.79$$

complete the square.

4/15

$$\propto \exp \left[ -\frac{1}{2} (Ax^2 + 2Bx + C) \right].$$

$$\propto \exp \left[ -\frac{1}{2} (Ax^2 + 2Bx) \right]. \text{ drop } C \propto$$

$$\propto \exp \left[ -\frac{1}{2} A \left( x^2 + 2\frac{B}{A}x \right) \right]$$

$$\propto \exp \left[ -\frac{1}{2} A \left( x^2 + 2\frac{B}{A}x + D \right) \right]. \text{ add } D \propto$$

$$\text{let } D = \left( \frac{B}{A} \right)^2$$

$$\propto \exp \left[ -\frac{1}{2} A \left( x - \frac{B}{A} \right)^2 \right].$$

$$\propto \exp \left[ -\frac{1}{2\sigma^2} (x - \mu)^2 \right].$$

$$\mu = \frac{B}{A}$$

$$\sigma^2 = \frac{1}{A}.$$

Example 8.1.2 Weighing and object

# Example 2, Problem 3.2

data

list(X = c(198.14, 198.45, 196.59, 197.64, 198.12), n = 5, sigma = 1) # frequentist estimate, mean = 197.79

inits

list(mu = 198)

model

model;

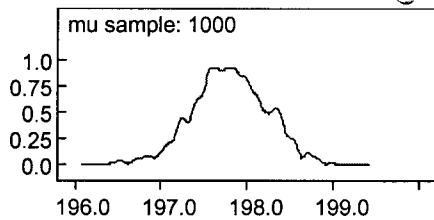
```

{
  for(i in 1:n){
    X[i] ~ dnorm(mu,tau)
  }
  mu ~ dnorm( 200.0,0.01) # prior mode = 200, 95% prior interval (180,220), Normal(200,10) prec = 1/10 = 0.1
  tau <- 1 / pow(sigma,2)
}

```

Output:

*posterior density of  $\mu$ .*



node	mean	sd	MC error	2.5%	median	97.5%	start	sample
mu	197.8	0.4371	0.01218	196.9	197.8	198.7	1001	1000

*Bayesian estimate of  $\mu$*

*Bayesian 95% credible interval for  $\mu$ .*

Example 8.1.3 Counting nice.

prior  $\lambda \sim \text{Gamma}(\alpha_0, k_0)$

$$p(\lambda) \propto \lambda^{\alpha_0-1} e^{-k_0 \lambda}$$

p. 95  $\alpha_0 = 4, k_0 = \frac{1}{3}$

likelihood  $x_1, \dots, x_n \stackrel{iid}{\sim} \text{Poi}(\lambda)$

p. 98 data  $n=50$

$$t = \sum x_i = 256.$$

$$p(x_1, \dots, x_n | \lambda) \propto \prod_{i=1}^n \lambda^{x_i} e^{-\lambda}$$

$$= \lambda^t e^{-n\lambda}$$

$$t = \sum x_i$$

posterior

$$p(\lambda | x_1, \dots, x_n) \propto \lambda^{\alpha_0-1} e^{-k_0 \lambda} \lambda^t e^{-n\lambda}$$

$$\propto \lambda^{\alpha_0+t-1} e^{-(k_0+n)\lambda}$$

$\therefore \lambda | x_1, \dots, x_n \sim \text{Gamma}(\alpha_n, k_n)$

$$\alpha_n = \alpha_0 + t$$

$$k_n = k_0 + n.$$

posterior mean.  $E[\lambda | x_1, \dots, x_n] = \int_0^\infty \lambda p(\lambda | x_1, \dots, x_n) d\lambda$

$$= \frac{\alpha_n}{k_n} = \frac{\alpha_0 + t}{k_0 + n}$$

p. 98

$$E[\lambda | x_1, \dots, x_n] = \frac{4 + 256}{\frac{1}{3} + 50} = 5.17$$

$$= \frac{\alpha_0}{k_0} \left( \frac{k_0}{k_0+n} \right) + \frac{t}{n} \left( \frac{n}{k_0+n} \right)$$



Example 8.1.3 Counting mice

# Example 3

data

list(T = 256, n = 50)

inits

list(lambda = 10)

model

model;

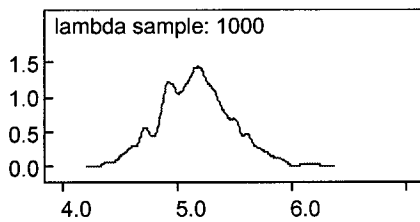
```

{
  T ~ dpois(lambda0)
  lambda ~ dgamma(4,0.3333)
}

```

Output:

*posterior of  $\lambda$*



node	mean	sd	MC error	2.5%	median	97.5%	start	sample
lambda	5.17	0.3148	0.01088	4.561	5.162	5.825	1001	1000

*bayesian estimate of  $\lambda$*

*bayesian 95% credible interval for  $\lambda$*

### Program 1.1 Power.odc

### Model B: Prior and likelihood coding

#### Model

```

{
  for (i in 1 : N) {
    lambda[i] ~ dgamma(alpha, beta)
    theta[i] <- lambda[i] * E[i]
    y[i] ~ dpois(theta[i])
  }
  alpha ~ dexp(1)
  beta ~ dgamma(0.1, 1.0)
}

```

#### B: Inits

```
list(alpha=1, beta=1)
```

#### A and B: data

```
list(E = c(94.3, 15.7, 62.9, 126, 5.24, 31.4, 1.05, 1.05, 2.1, 10.5),
     y = c( 5, 1, 5, 14, 3, 19, 1, 1, 4, 22), N = 10)
```

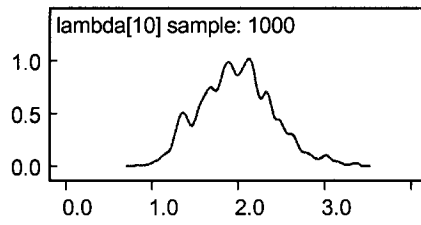
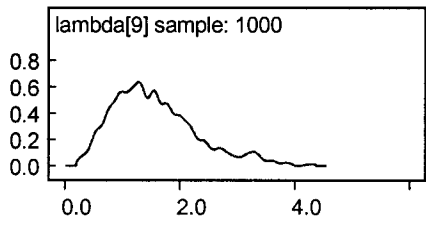
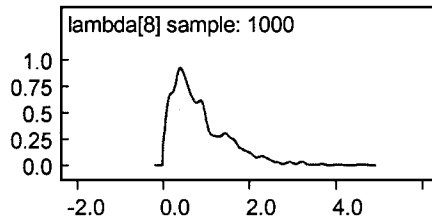
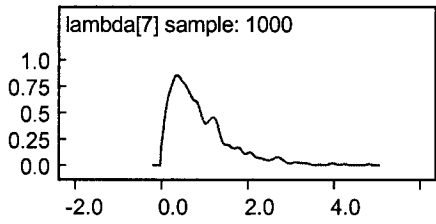
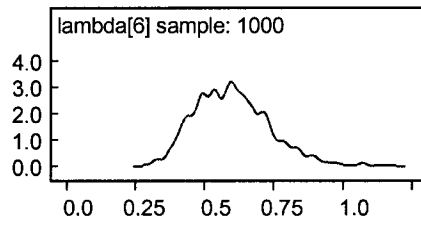
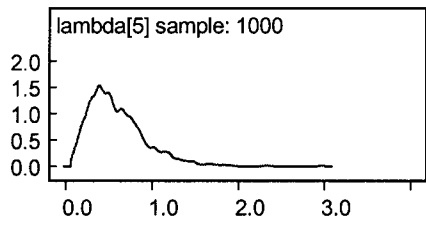
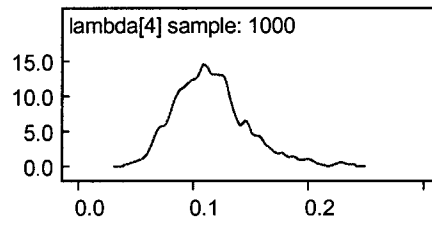
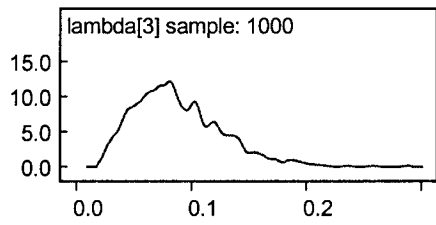
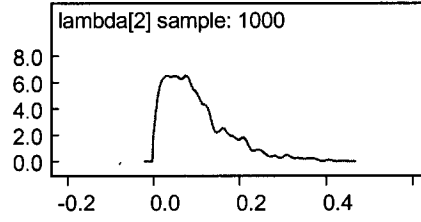
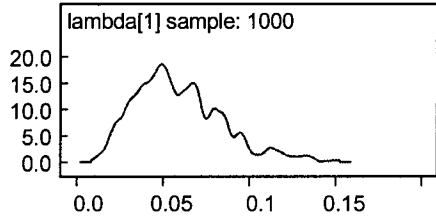
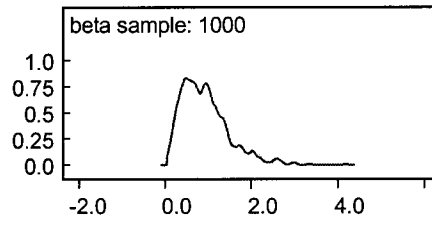
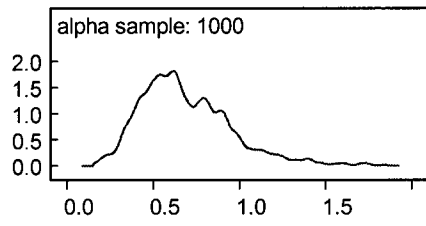
#### Output:

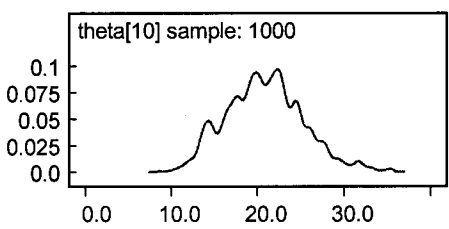
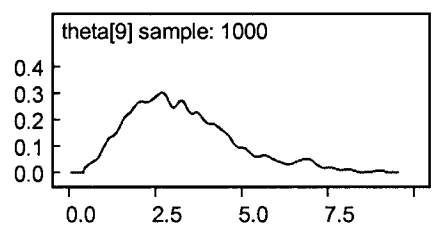
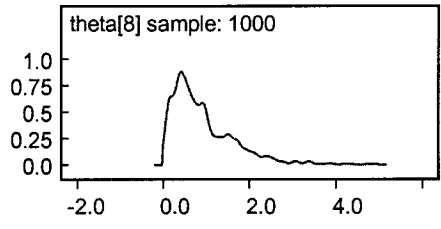
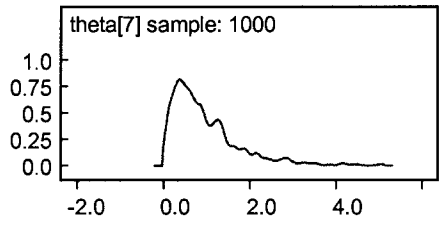
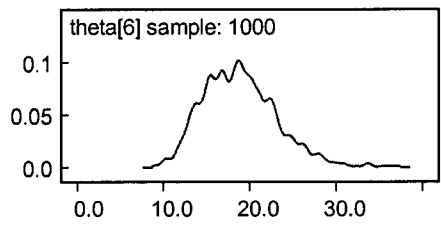
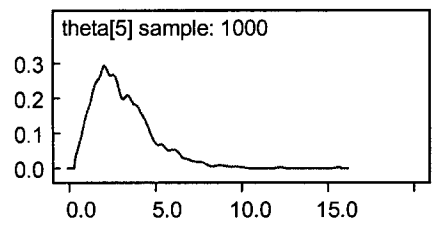
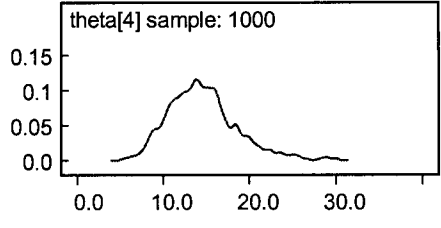
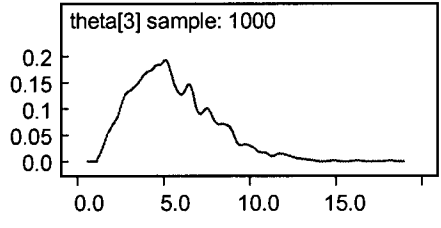
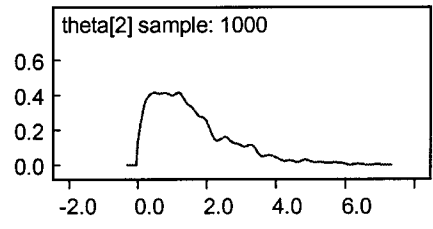
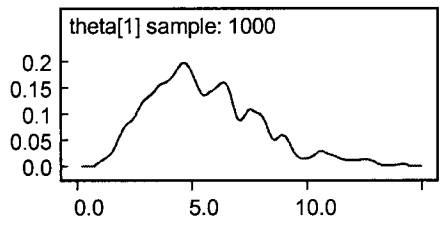
node	mean	sd	MC error	2.5%	median	97.5%	start	sample
alpha	0.6922	0.2642	0.01444	0.2986	0.6442	1.342	1001	1000
beta	0.9301	0.5498	0.02954	0.1864	0.8518	2.28	1001	1000
lambda[1]	0.05928	0.02521	8.073E-4	0.0206	0.05517	0.1197	1001	1000
lambda[2]	0.09994	0.07535	0.002759	0.007173	0.08192	0.3051	1001	1000
lambda[3]	0.08868	0.03812	0.001307	0.03056	0.08325	0.1804	1001	1000
lambda[4]	0.1161	0.03169	9.349E-4	0.06509	0.1129	0.1904	1001	1000
lambda[5]	0.5974	0.3333	0.01431	0.1458	0.5296	1.398	1001	1000
lambda[6]	0.6026	0.1363	0.004635	0.3825	0.5936	0.9015	1001	1000
lambda[7]	0.9059	0.7322	0.02253	0.08242	0.7118	2.832	1001	1000
lambda[8]	0.8767	0.6935	0.02714	0.07072	0.6983	2.63	1001	1000
lambda[9]	1.578	0.7369	0.02195	0.4871	1.469	3.351	1001	1000
lambda[10]	1.989	0.4236	0.0124	1.249	1.981	2.916	1001	1000
theta[1]	5.59	2.377	0.07613	1.942	5.203	11.29	1001	1000
theta[2]	1.569	1.183	0.04332	0.1126	1.286	4.789	1001	1000
theta[3]	5.578	2.398	0.08219	1.922	5.236	11.35	1001	1000
theta[4]	14.63	3.993	0.1178	8.201	14.22	23.99	1001	1000
theta[5]	3.131	1.746	0.07498	0.7639	2.775	7.324	1001	1000
theta[6]	18.92	4.281	0.1455	12.01	18.64	28.31	1001	1000
theta[7]	0.9512	0.7688	0.02365	0.08654	0.7474	2.973	1001	1000
theta[8]	0.9205	0.7282	0.02849	0.07426	0.7332	2.762	1001	1000
theta[9]	3.314	1.547	0.0461	1.023	3.084	7.038	1001	1000
theta[10]	20.88	4.448	0.1302	13.12	20.8	30.62	1001	1000

↑  
 parameter  
 estimates

↑  
 credible  
 intervals

11/5





# Program 1.2 Remission data.odc

## # Prior and Likelihood Model

```

model {
  for (i in 1:42) {
    t[i] ~ dnorm(mu[Tr[i]], tau[Tr[i]]) l(min[i], )
  }
  for (j in 1:2) {
    tau.mu[j] <- n0*tau[j]
    mu[j] ~ dnorm(0,tau.mu[j])
  }
  # independent priors
  # mu[j] ~ dnorm(0,0.001)
  tau[j] ~ dgamma(1,0.001)
  sigma2[j] <- 1/tau[j]
  sigma[j] <- sqrt(sigma2[j])
}

```

## # Data

```

list(n0=1,t=c(NA,6,6,6,7,NA,NA,10,NA,13,16,NA,NA,NA,22,23,NA,NA,NA,
NA,NA,1,1,2,2,3,4,4,5,5,8,8,8,8,11,11,12,12,15,17,22,23),
min=c(6,0,0,0,0,9,10,0,11,0,0,17,19,20,0,0,25,32,32,34,35,0,0,
0,0,0,0,0,0,0,0,0,0,0,0,0,0,0,0,0,0,0,0),
Tr=c(1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,2,2,2,2,2,2,2,2,2,2,2,2,2,2,
2,2))

```

## # Inits

```
list(mu=c(10,10),tau=c(1,1))
```

*A survival time is recorded if the subject leaves the study before the cancer returns.*

## Output:

	node	mean	sd	MC error	2.5%	median	97.5%	start	sample
→	mu[1]	24.61	4.743	0.3027	17.02	23.98	36.23	1001	1000
	mu[2]	8.282	1.388	0.0438	5.564	8.288	11.25	1001	1000
	sigma[1]	16.12	4.18	0.273	10.04	15.41	27.05	1001	1000
	sigma[2]	6.478	0.9795	0.03307	4.804	6.429	8.628	1001	1000
	sigma2[1]	277.3	156.1	10.36	100.7	237.5	732.0	1001	1000
	sigma2[2]	42.93	13.46	0.4505	23.08	41.34	74.44	1001	1000
	t[1]	28.89	14.9	0.7895	8.026	26.32	65.14	1001	1000
	t[6]	29.47	13.1	0.446	10.69	27.4	59.48	1001	1000
	t[7]	30.04	13.9	0.6972	11.1	27.7	63.34	1001	1000
	t[9]	30.6	13.84	0.5875	12.25	28.3	62.58	1001	1000
	t[12]	33.12	12.0	0.5798	17.6	30.6	63.38	1001	1000
	t[13]	33.97	11.78	0.4489	19.77	30.83	64.41	1001	1000
	t[14]	35.0	12.32	0.4954	20.49	32.28	67.19	1001	1000
	t[17]	37.54	11.04	0.444	25.45	34.7	65.17	1001	1000
	t[18]	42.47	9.801	0.4519	32.38	39.53	71.21	1001	1000
	t[19]	42.75	9.926	0.4332	32.31	39.68	68.39	1001	1000
	t[20]	44.09	9.785	0.4516	34.29	41.47	68.46	1001	1000
	t[21]	45.16	9.496	0.4357	35.27	42.3	70.68	1001	1000
	tau[1]	0.004592	0.002198	1.234E-4	0.001377	0.004216	0.01006	1001	1000
	tau[2]	0.02544	0.007504	2.579E-4	0.01344	0.02421	0.04348	1001	1000
	tau.mu[1]	0.004592	0.002198	1.234E-4	0.001377	0.004216	0.01006	1001	1000
	tau.mu[2]	0.02544	0.007504	2.579E-4	0.01344	0.02421	0.04348	1001	1000

*treatment group 1.*

*logistic  
variables*

↑  
*credible  
intervals*

14/15

