Hierarchical Bayesian model for certification of a country as "free" from an animal disease

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Outline:

- Background
- Disease Freedom
- Current Approaches
 - Qualitative
 - Quantitative
- Hierarchical Model
- Bayesian Approach: Gibbs Sampling
- Example: Newcastle Disease

Background:

- Risk analysis for trade in animal products.
- Countries are interested in new trading opportunities and maintaining current trade.
- Is the country disease free?
- Needed for risk analysis to make policy decisions.

Disease Freedom:

- "Disease freedom"
 - requires a *perfectly* sensitive test
 - All animals
 - All negative tests
- Level of Disease Freedom may not mean total freedom in all situations.
 - prevalence < threshold

Current Approaches: • Qualitative • Quantitative - Monte Carlo Simulation

Data: • Two-stage cluster sample. ullet Random selection of k herds and a random sample of n animals within each herd.

Bayesian Model:

Country level:

Z = "true" status of the country.

$$Z = 1$$
 α diseased $= 0 \quad 1 - \alpha$

Herd level: (latent data)

 $t_i =$ "true" status of herd i.

$$t_i = 1 \quad \tau \qquad i^{th} \text{ herd is diseased}$$

$$=$$
 0 $1-\tau$

Within herd level:

Assume two populations exist from which the i^{th} herd can be selected, either diseased or non-diseased.

 λ_i = prevalence of the disease in the population from which the i^{th} herd was sampled.

 π_i = prevalence of the disease in a diseased herd.

Prevalence Within Diseased Herds:

$$\pi_i \sim beta(a,b)$$

where (a, b) is random.

Prevalence Within Herds:

$$\lambda_i | \pi_i = \pi_i \quad \tau$$

$$= 0 \quad 1 - \tau$$

So

$$\lambda_i | \pi_i, \tau \sim \pi_i Bernoulli(\tau)$$

Within herd level: (latent data)

 v_{ij} = "true" status of animal j within herd i.

$$v_{ij} = 1$$
 j^{th} animal in i^{th} herd is diseased
= 0 not diseased

So

$$v_{ij}|Z=1, t_i=1 \sim Bernoulli(\pi_i)$$

Data:

 $X_{ij} = 1$ if the j^{th} animal in the i^{th} herd tests positive = 0 otherwise test negative

Imperfect diagnostic test:

- 1. sensitivity $\eta = P(T^+|D) < 1$
- 2. specificity $\theta = P(T^-|\bar{D}) < 1$

 $X_i = \sum_j X_{ij} = \text{number of test positive animals in the } i^{th} \text{ herd.}$

If Z=1

$$X_i|Z=1, \lambda_i, \eta, \theta \sim Bin[n_i, \lambda_i \eta + (1-\lambda_i)(1-\theta)]$$

If Z = 0

$$X_i|Z=0 \sim Bin[n_i,(1-\theta)]$$

Important: This model generates correlation in disease status between animals within each herd, but leaves the disease status independent between herds.

Bayesian Statistics:

X is the data vector and Θ is the parameter set.

Model: (or likelihood)

$$p(\mathbf{X}|\mathbf{\Theta})$$

Prior distribution:

$$p(\mathbf{\Theta})$$

Joint distribution:

$$p(\mathbf{X}, \mathbf{\Theta}) = p(\mathbf{X}|\mathbf{\Theta})p(\mathbf{\Theta})$$

Posterior distribution: (Bayes Theorem)

$$p(\mathbf{\Theta}|\mathbf{X}) = \frac{p(\mathbf{X}|\mathbf{\Theta})p(\mathbf{\Theta})}{\int p(\mathbf{X}|\mathbf{\Theta})p(\mathbf{\Theta})d\mathbf{\Theta}}$$

$$\propto p(\mathbf{X}|\mathbf{\Theta})p(\mathbf{\Theta})$$

Gibbs Sampler: (ref. Gelfand and Smith 1990)

Two parameter case: $\mathbf{\Theta} = (\theta_1, \theta_2) \ p(\mathbf{\Theta}|\mathbf{X}) = p(\theta_1, \theta_2|\mathbf{X})$

Want the marginals, $p(\theta_1|\mathbf{X})$ and $p(\theta_2|\mathbf{X})$

Can get the conditional marginals, $p(\theta_1|\mathbf{X}, \theta_2)$ and $p(\theta_2|\mathbf{X}, \theta_1)$

Initial value $(\theta_1^{(0)}, \theta_2^{(0)})$ for h = 1, ..., Reps

- 1. Sample $\theta_1^{(h)}$ from $p(\theta_1|\mathbf{X}, \theta_2^{(h-1)})$.
- 2. Sample $\theta_2^{(h)}$ from $p(\theta_2|\mathbf{X}, \theta_1^{(h)})$.
- 3. Set h = h + 1 and go to 1.

Inferences are calculated using the simulated data after the chain has stabilized.

$$(\theta_1^{(BurnIn+1)}, \theta_2^{(BurnIn+1)}), ..., (\theta_1^{(Reps)}, \theta_2^{(Reps)})$$

The theory behind the Gibbs sampler says, $\boldsymbol{\Theta}^{(1)},...,\boldsymbol{\Theta}^{(Reps)}$ are realizations of a stationary Markov Chain, with transition probability from $\boldsymbol{\Theta}^{(h-1)}$ to $\boldsymbol{\Theta}^{(h)}$,

$$T(\mathbf{\Theta}^{(h-1)}, \mathbf{\Theta}^{(h)}) = p(\theta_1 | \mathbf{X}, \theta_2^{(h-1)}) p(\theta_2 | \mathbf{X}, \theta_1^{(h)})$$

Justified by Ergodic theory, we can calculate estimates of the parameter using means. For example,

$$\hat{\theta}_1 = \frac{1}{Reps} \sum_{h} \theta_1^{(i)} \xrightarrow{a.s.} E[\theta_1 | \mathbf{X}],$$

$$Reps \rightarrow \infty$$
.

Priors:

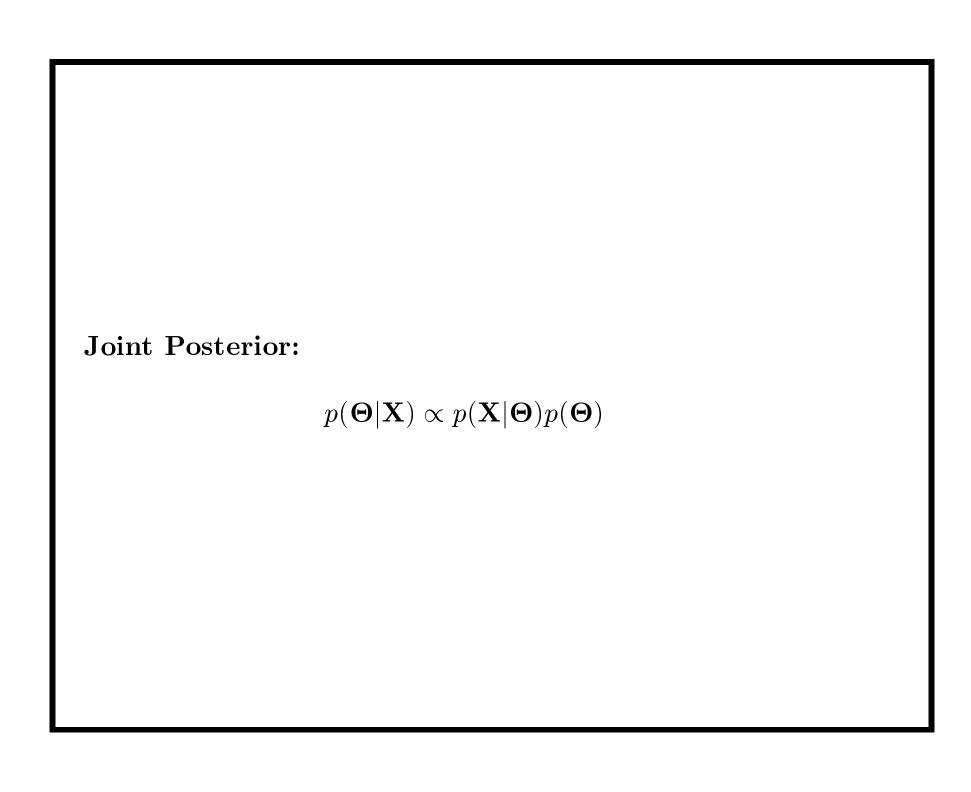
- Country level, $\alpha \sim beta(a_{\alpha}, b_{\alpha})$.
- Herd level, $\tau \sim beta(a_{\tau}, b_{\tau})$.
- Within herd level, $\pi_i \sim beta(a, b)$.
- The distribution from which the prevalence a diseased herd results from is assumed to be a beta distribution with unknown parameters (a, b). Adaptive Rejection Sampling.
- Sensitivity, $\eta \sim beta(a_{\eta}, b_{\eta})$.
- Specificity, $\theta \sim beta(a_{\theta}, b_{\theta})$.

The Likelihood:

$$X_{ij}|Z=1,\{v_{ij}\},\{t_i\},\{\pi_i\},\eta,\theta \sim Bernoulli[\eta^{v_{ij}}(1-\theta)^{1-v_{ij}}].$$

$$X_{ij}|Z=0 \sim Bernoulli[(1-\theta)].$$

Parameter Set: $\Theta = \{Z, \{v_{ij}\}, \{t_i\}, \{\pi_i\}, (a, b), \alpha, \tau, \eta, \theta\}$



Conditional Marginal Posterior Distributions:

• "true" status of each animal.

$$v_{ij}|Z = 1, X_{ij} = 1, t_i = 1, "rest" \sim Bernoulli(P(D|X_{ij} = 1))$$

$$P(D|X_{ij} = 1) = \frac{\eta \pi_i}{\eta \pi_i + (1 - \theta)(1 - \pi_i)}$$

$$v_{ij}|Z = 1, X_{ij} = 0, t_i = 1, "rest" \sim Bernoulli(P(D|X_{ij} = 0))$$

$$P(D|X_{ij} = 0) = \frac{(1 - \eta)\pi_i}{(1 - \eta)\pi_i + \theta(1 - \pi_i)}$$

$$v_{ij}|Z=1, t_i=0, "rest" \sim Bernoulli(0)$$

• "true" status of each herd

If $v_{ij} = 1$ for any j within herd i, then

$$t_i|\{X_{ij}\},$$
 "rest" $\sim Bernoulli(1)$.

If $v_{ij} = 0$ for all j within herd i, then

$$t_i | \{X_{ij}\}, "rest" \sim Bernoulli\left(\frac{(1-\pi)^n \tau}{(1-\pi_i)^n + 1(1-\tau)}\right).$$

If Z=0, then

$$t_i|Z=0$$
, "rest" $\sim Bernoulli(0)$.

• prevalence of a diseased herd.

$$\pi_i|\{X_{ij}\}, "rest" \sim beta\left(a + \sum_j v_{ij}, b + n - \sum_j v_{ij}\right)$$

• hyperparameters for the distribution on the π_i .

 $(a,b)|\{X_{ij}\},$ "rest" sampled using Adaptive Rejection.

• probability the country is diseased.

$$\alpha | \{X_{ij}\}, "rest" \sim beta (a_{\alpha} + Z, b_{\alpha} + (1 - Z))$$

• probability a herd is diseased.

$$\tau | \{X_{ij}\}, "rest" \sim beta\left(a_{\tau} + \sum_{i} t_{i}, b_{\tau} + k - \sum_{i} t_{i}\right)$$

• sensitivity.

$$\eta | \{X_{ij}\}, "rest" \sim beta \left(a_{\eta} + \sum x_{ij}v_{ij}, b_{\eta} + \sum (1 - x_{ij})v_{ij}\right)$$

• specificity.

$$\theta | \{X_{ij}\}, \text{"rest"} \sim beta(a_{\theta} + \sum_{i=1}^{n} (1 - x_{ij})(1 - v_{ij}), b_{\theta} + \sum_{i=1}^{n} x_{ij}(1 - v_{ij}))$$

• country status.

$$Z|\{X_{ij}\}, "rest" \sim Bernoulli\left(\frac{\alpha(1-\tau)^k}{\alpha(1-\tau)^k+(1-\alpha)}\right)$$

Adaptive Rejection Sampling: (ref. Gilks and Wild 1992)

$$\prod_{i=1}^{k} \left[\frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \pi_i^{a-1} (1-\pi_i)^{b-1} p(a) \right]^{t_i}$$

log-concavity in a given b.

Steps To Perform The Gibbs Sampler:

Given the initial values:

1. $v_{ij}^{(h)} \sim Bernoulli, 1, ..., k, j = 1, ..., n.$

2. $t_i^{(h)} \sim Bernoulli, i = 1, ..., k$.

3. sample (a, b) using Adaptive Rejection.

4. $\pi_i^{(h)} \sim beta, i = 1, ..., k.$

5. $\alpha^{(h)} \sim beta$.

6. $\tau^{(h)} \sim beta$.

7. $\eta^{(h)} \sim beta$.

8. $\theta^{(h)} \sim beta$.

9. $Z^{(h)} \sim Bernoulli$.

Conclusions:

- Bayesian approach.
 - Uses prior knowledge.
 - Latent data.
 - Imperfect test.
- Variable within herd-level prevalence.
- Country/Herd/Animal inference.
- Priors for other risk analysis.

Further Work: • Variable sample size within herds. • Application to continuous surveillance.

References:

Audigé, L., Beckett (1999) A Quantitative Assessment of the Validity of Animal-health Surveys Using Stochastic Modeling *Journal of Preventitive Veterinary Medicine* **38**, 259-276.

Gelfand, A.E. and Smith, A.F.M. (1990) Sampling-Based Approaches to Calculating Marginal Densities *Journal of the American Statistical Association* **85**, 398-409.

Gilks, W.R. and Wild, P. (1992)Adaptive Rejection Sampling for Gibbs Sampling. *Applied Statistics* **41**, 337-348.

Gohm, Thür, Audigé, and Hoffmann (1999) A Survey of Newcastle Disease in Swiss Lay-hen Flocks Using Serological Testing and Simulation Modeling. *Journal of Preventitive Veterinary Medicine* **38**, 277-288.

Vose, D.J. (1998) Risk Analysis in Relation to the Importation and Exportation of Animal Products. Rev. sci. tech. Off. int. Epiz 16, 17-29.